

Supplementary Appendix to LEARNING UNDER AMBIGUITY*

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Abstract

This appendix provides a detailed treatment of the portfolio choice problem studied in our paper “Learning under Ambiguity”, filling in details of calculations that were omitted in the text of the paper.

1 Investor Problem

Section 1 of this appendix restates the portfolio problem studied in Section 5 of “Learning under Ambiguity”. Section 2 defines beliefs and derives the dynamics of beliefs under ambiguity. Section 3 computes optimal portfolio weights for a benchmark Bayesian investor and a myopic ambiguity-averse investor.

Time is measured in months; there are k trading dates per month. The state space is $S = \{1, 0\}$.¹ The return on stocks $R(s_t) = e^{r(s_t)}$ realized in period t is either high or low: we fix (log) return realizations $r(1) = \sigma/\sqrt{k}$ and $r(0) = -\sigma/\sqrt{k}$. In addition to stocks, the investor also has access to a riskless asset with constant per period interest rate $R^f = e^{r^f/k}$, where $r^f < \sigma$. We consider investors who plan for T months starting in month t and who care about terminal wealth according to the utility function $V_T(W_{t+T}) = \log W_{t+T}$. Investors may rebalance their portfolio at all $k(T-t)$ trading dates between t and T . Let $\omega_{t,T,k}^*$ denote the optimal fraction of wealth invested in stocks at date t for an investor who plans for T months, when there are k trading dates per month.

Consider the investor’s problem when beliefs are given by a general process of one-step-ahead conditionals $\{\mathcal{P}_\tau(s^\tau)\}$. The history s^τ of state realizations up to trading date

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¹In this appendix, we call the binary states 1 and 0, rather than *hi* and *lo* as in the text of “Learning under Ambiguity”, because this allows simpler notation below.

$\tau = t + j/k$ – the j th trading date in month $t + 1$ – can be summarized by the fraction ϕ_τ of the state $s_t = 1$ observed up to τ . The value function of the log investor takes the form $V_\tau(W_\tau, s^\tau) = h_\tau(\phi_\tau) + \log W_\tau$. The process $\{h_\tau\}$ satisfies $h_{t+T} = 0$ and

$$\begin{aligned} h_\tau(\phi_\tau) &= \max_{\omega_\tau} \min_{p_\tau \in \mathcal{P}_\tau(s^\tau)} E^{p_\tau} [\log(R^f + (R(s_{\tau+1/k}) - R^f)\omega_\tau) + h_{\tau+1/k}(\phi_{\tau+1/k})] \\ &= \min_{p_\tau \in \mathcal{P}_\tau(s^\tau)} \max_{\omega_\tau} E^{p_\tau} [\log(R^f + (R(s_{\tau+1/k}) - R^f)\omega_\tau) + h_{\tau+1/k}(\phi_{\tau+1/k})], \end{aligned} \quad (1)$$

where we have used the minimax theorem to reverse the order of optimization.

We are interested in the optimal portfolio of an investor who has seen a monthly sample of log real US stock returns $\{r_\tau\}_{\tau=1}^t$. However, agents in the model are assumed to observe not only monthly returns, but actually binary returns $R(s_\tau)$ at every trading date. To model an agent’s history up to date t , we thus construct a sample of tk realizations of $R(s_\tau)$ such that the implied empirical distribution of monthly log returns is the same as that in the data. Let σ denote the monthly standard deviation of log returns. The sample of binary returns up to some integer date τ is summarized by the fraction of states $s_t = 1$, defined by

$$\phi_\tau = \hat{\phi}_k(\bar{r}_\tau) := \frac{1}{2} + \frac{1}{2} \frac{\bar{r}_\tau}{\sigma\sqrt{k}}. \quad (2)$$

where $\bar{r}_\tau := \frac{1}{\tau} \sum_{j=1}^\tau r_j$ is the mean of the monthly return sample. For given k , the sequence $\{\phi_\tau\}$ pins down a sequence of monthly log returns $\{\sum_{j=1}^k \log R(s_{\tau+j/k})\}$ that is identical to the data sample $\{r_\tau\}$.

2 Beliefs

As a Bayesian benchmark, we assume that the investor has an improper beta prior over the probability p of the high state, so that the posterior mean of p after t months (or tk state realizations) is equal to ϕ_τ , the maximum likelihood estimator of p . The Bayesian’s probability of a high state next period is then also given by ϕ_τ . The optimal portfolio follows from solving (1) when $\mathcal{P}_\tau(s^\tau)$ is a singleton set containing only the measure that puts probability ϕ_τ on the high state.

For an ambiguity-averse investor, beliefs are defined as in Section 3 of “Learning under Ambiguity”. We briefly review the general model here. Beliefs are represented by

$$(\Theta, \mathcal{M}_0, \mathcal{L}_k, \alpha),$$

where Θ is a parameter space, \mathcal{M}_0 is a set of priors on Θ , \mathcal{L}_k is a set of likelihoods and α is a number between zero and one. A *theory* is a pair (μ_0, ℓ^t) , where μ_0 is a prior belief on Θ and $\ell^t = (\ell_{1/k}, \ell_{2/k}, \dots, \ell_t) \in \mathcal{L}_k^{tk}$ is a *sequence* of likelihoods. Let $\mu_t(\cdot; s^t, \mu_0, \ell^t)$ denote the posterior derived from the theory (μ_0, ℓ^t) by Bayes’ Rule, given the data s^t .

The set of posteriors contains posteriors that are based on theories not rejected by a

likelihood ratio test:

$$\begin{aligned} \mathcal{M}_{t,k}^\alpha(s^t) = \{ & \mu_t(s^t; \mu_0, \ell^t) : \mu_0 \in \mathcal{M}, \ell^t \in \mathcal{L}_k^t, \\ & \int \prod_{j=1}^t \ell_j(s_j | \theta) d\mu_0(\theta) \geq \alpha \max_{\substack{\tilde{\mu}_0 \in \mathcal{M}_0 \\ \tilde{\ell}^t \in \mathcal{L}_k^t}} \int \prod_{j=1}^t \tilde{\ell}_j(s_j | \theta) d\tilde{\mu}_0\}. \end{aligned} \quad (3)$$

The set of one-step-ahead conditional belief is defined by

$$\mathcal{P}_t(s^t) = \left\{ p_t(\cdot) = \int_{\Theta} \ell(\cdot | \theta) d\mu_t(\theta) : \mu_t \in \mathcal{M}_t^\alpha(s^t), \ell \in \mathcal{L}_k \right\}, \quad (4)$$

This process enters the specification of recursive multiple priors preferences in (1).

The particular representation $(\Theta, \mathcal{M}_0, \mathcal{L}_k, \alpha)$ assumed for the portfolio choice problem is defined as follows. The ambiguity averse investor perceives the mean monthly log return as $\theta + \lambda_t$, where $\theta \in \Theta := \mathbb{R}$ is fixed and can be learned, while λ_t is driven by many poorly understood factors affecting returns and can *never* be learned. The set \mathcal{L}_k consists of all $\ell(\cdot | \theta)$ such that

$$\ell^k(hi|\theta) = \frac{1}{2} + \frac{1}{2} \frac{\theta + \lambda}{\sigma\sqrt{k}}, \quad \text{for some } \lambda \text{ with } |\lambda| < \bar{\lambda}. \quad (5)$$

The set of priors \mathcal{M}_0 on Θ consists of Dirac measures. For simplicity, we write $\theta \in \mathcal{M}_0$ if the Dirac measure on θ is included in the set of priors, and we define

$$\mathcal{M}_0 = \{\theta : |\theta| \leq \bar{\lambda} + 1/\sigma\}.$$

The condition ensures that the probability (5) remains between zero and one for all $k \geq 1$.

Beliefs depend on history only via the fraction of high returns $\phi_t = \hat{\phi}_k(\bar{r}_t)$. We write $\theta \in \mathcal{M}_{t,k}^\alpha(\bar{r}_t)$ if the Dirac measure on θ is included in the posterior set at the end of month t after history $\hat{\phi}_k(\bar{r}_t)$. The posterior set for fixed k can be characterized as follows:

Proposition S1. *The posterior set is a subinterval $[\underline{\theta}_k(\bar{r}_t), \bar{\theta}_k(\bar{r}_t)]$ of \mathcal{M}_0 with both bounds strictly increasing in \bar{r}_t .*

Proof. The history s^t consists of tk realizations of the state. Write the likelihood of a sample s^t under some theory, here identified with a pair (θ, λ^t) , as

$$L_k(s^t, \theta, \lambda^t) = \prod_{\tau=0}^{t-1} \prod_{j=1}^k \left(\frac{1}{2} + \frac{1}{2} \frac{\theta + \lambda_{\tau+j/k}}{\sigma\sqrt{k}} \right)^{s_{\tau+j/k}} \left(\frac{1}{2} - \frac{1}{2} \frac{\theta + \lambda_{\tau+j/k}}{\sigma\sqrt{k}} \right)^{1-s_{\tau+j/k}}. \quad (6)$$

Let $\tilde{\lambda}^t$ denote the sequence that maximizes (6) for fixed θ . This sequence is independent of θ and has $\tilde{\lambda}_i = \bar{\lambda}$ if $s_i = 1$ and $\tilde{\lambda}_i = -\bar{\lambda}$ if $s_i = 0$, for all $i \leq t$. It follows that

$L_k(s^t, \theta, \tilde{\lambda}^t)$ depends on the sample only through the fraction ϕ_t of high returns observed. The posterior set can be written

$$\mathcal{M}_{t,k}^\alpha(s^t) = \left\{ \theta : \frac{1}{tk} \log L(s^t, \theta, \tilde{\lambda}^t) \geq \max_{\tilde{\theta}} \frac{1}{tk} \log L(s^t, \tilde{\theta}, \tilde{\lambda}^t) - \frac{1}{tk} \log(\alpha) \right\} \quad (7)$$

Indeed, if $\theta \in \mathcal{M}_{t,k}^\alpha$, then there exists *some* λ^t such that the theory (θ, λ^t) passes the criterion for an admissible theory put forward in (3). Thus the theory $(\theta, \tilde{\lambda}^t)$ must also pass that criterion, since its likelihood is at least as high. In contrast, if $\theta \notin \mathcal{M}_{t,k}^\alpha$, then there is *no* λ^t such that the theory (θ, λ^t) passes the criterion.

Using the fraction of high states $\phi_t = \frac{1}{tk} \sum_i s_i$, rewrite the log data density as

$$\log L(s^t, \theta, \tilde{\lambda}^t) = f(\theta; \phi_t) := \phi_t \log \left(\frac{1}{2} + \frac{1}{2} \frac{\theta + \bar{\lambda}}{\sigma \sqrt{k}} \right) + (1 - \phi_t) \log \left(\frac{1}{2} - \frac{1}{2} \frac{\theta - \bar{\lambda}}{\sigma \sqrt{k}} \right).$$

The likelihood ratio criterion becomes

$$f(\theta; \phi_t) \geq \max_{\tilde{\theta}} \{f(\tilde{\theta}; \phi_t)\} - \frac{1}{tk} \log(\alpha). \quad (8)$$

For $\phi_t \in (0, 1)$, the function f is strictly concave and achieves a unique maximum at the MLE

$$\theta^*(\phi_t) = (2\phi_t - 1) (\sigma \sqrt{k} + \bar{\lambda}) = (2\phi_k(\bar{r}_t) - 1) (\sigma \sqrt{k} + \bar{\lambda}) = \bar{r}_t \left(1 + \bar{\lambda} / \sigma \sqrt{k} \right).$$

Since f is strictly concave and $\lim_{\theta \rightarrow \bar{\lambda} + 1/\sigma} f(\theta; \phi_t) = \lim_{\theta \rightarrow -\bar{\lambda} - 1/\sigma} f(\theta; \phi_t) = -\infty$, the set of θ s that pass the likelihood ratio test is a subinterval of \mathcal{M}_0 , with bounds that satisfy (8) with equality. Using $\phi_t = \phi_k(\bar{r}_t)$, the posterior set can be written as an interval $[\underline{\theta}_k(\bar{r}_t), \bar{\theta}_k(\bar{r}_t)]$.

To see why both bounds are strictly increasing in \bar{r}_t , suppose that $\tilde{\theta}$ satisfies (8) with equality. Apply the implicit function theorem to obtain

$$\left. \frac{d\theta}{d\phi} \right|_{\theta=\tilde{\theta}} = \frac{-f_2(\tilde{\theta}; \phi_t) + f_2(\theta^*(\phi_t); \phi_t)}{f_1(\tilde{\theta}; \phi_t)},$$

where f_i is the derivative of f with respect to its i th argument. Since f is strictly concave, $f_1(\tilde{\theta}; \phi_t) > 0$ if $\tilde{\theta} < \theta^*(\phi_t)$ and $f_1(\tilde{\theta}; \phi_t) < 0$ if $\tilde{\theta} > \theta^*(\phi_t)$. In addition, it can be verified that $f_{21}(\tilde{\theta}; \phi_t) > 0$, so that $f_2(\tilde{\theta}; \phi_t) - f_2(\theta^*(\phi_t); \phi_t) < 0$ if $\tilde{\theta} < \theta^*(\phi_t)$, but $f_2(\tilde{\theta}; \phi_t) - f_2(\theta^*(\phi_t); \phi_t) > 0$ if $\tilde{\theta} > \theta^*$. Taken together, these facts imply that $d\theta/d\phi > 0$. Since $\phi'_k(\bar{r}_t) > 0$, it follows that the bounds are strictly increasing in \bar{r}_t . ■

A simple formula for the posterior set is obtained by taking the limit as $k \rightarrow \infty$:

Proposition S2. *The limit of the posterior set is given by*

$$\left[\lim_{k \rightarrow \infty} \underline{\theta}_k(\bar{r}_t), \lim_{k \rightarrow \infty} \bar{\theta}_k(\bar{r}_t) \right] = \left[\bar{r}_t - t^{-\frac{1}{2}} \sigma b_\alpha, \bar{r}_t + t^{-\frac{1}{2}} \sigma b_\alpha \right], \quad (9)$$

where $b_\alpha = \sqrt{-2 \log \alpha}$.

Proof. Substitute for θ^* and $\phi_t = \phi_k(\bar{r}_t)$ defined in (2) to obtain

$$\begin{aligned} & \left(1 + \frac{\bar{r}_t}{\sigma\sqrt{k}} \right) \log \left(1 + \frac{\theta + \bar{\lambda}}{\sigma\sqrt{k}} \right) + \left(1 - \frac{\bar{r}_t}{\sigma\sqrt{k}} \right) \log \left(1 - \frac{\theta - \bar{\lambda}}{\sigma\sqrt{k}} \right) \\ & \geq \left(1 + \frac{\bar{r}_t}{\sigma\sqrt{k}} \right) \log \left(1 + \frac{\bar{r}_t + \bar{\lambda}}{\sigma\sqrt{k}} + \frac{\bar{r}_t \bar{\lambda}}{\sigma^2 k} \right) \\ & \quad + \left(1 - \frac{\bar{r}_t}{\sigma\sqrt{k}} \right) \log \left(1 - \frac{\bar{r}_t - \bar{\lambda}}{\sigma\sqrt{k}} - \frac{\bar{r}_t \bar{\lambda}}{\sigma^2 k} \right) - \frac{2}{tk} \log(\alpha) \end{aligned}$$

For each term of the form $\log(1+x)$, perform a Taylor expansion around $x=0$:

$$\begin{aligned} & \left(1 + \frac{\bar{r}_t}{\sigma\sqrt{k}} \right) \left(\frac{\theta + \bar{\lambda}}{\sigma\sqrt{k}} - \frac{1}{2} \frac{(\theta + \bar{\lambda})^2}{\sigma^2 k} + O(k^{-\frac{3}{2}}) \right) \\ & + \left(1 - \frac{\bar{r}_t}{\sigma\sqrt{k}} \right) \left(-\frac{\theta - \bar{\lambda}}{\sigma\sqrt{k}} - \frac{1}{2} \frac{(\theta - \bar{\lambda})^2}{\sigma^2 k} + O(k^{-\frac{3}{2}}) \right) \\ & \geq \left(1 + \frac{\bar{r}_t}{\sigma\sqrt{k}} \right) \left(\frac{\bar{r}_t + \bar{\lambda}}{\sigma\sqrt{k}} - \frac{1}{2} \frac{(\bar{r}_t + \bar{\lambda})^2}{\sigma^2 k} + \frac{\bar{r}_t \bar{\lambda}}{\sigma^2 k} + O(k^{-\frac{3}{2}}) \right) \\ & + \left(1 - \frac{\bar{r}_t}{\sigma\sqrt{k}} \right) \log \left(-\frac{\bar{r}_t - \bar{\lambda}}{\sigma\sqrt{k}} - \frac{1}{2} \frac{(\bar{r}_t - \bar{\lambda})^2}{\sigma^2 k} - \frac{\bar{r}_t \bar{\lambda}}{\sigma^2 k} + O(k^{-\frac{3}{2}}) \right) - \frac{2}{tk} \log(\alpha) \end{aligned}$$

Multiply out to obtain

$$\begin{aligned} & \frac{\bar{\lambda}}{\sigma\sqrt{k}} + \frac{2\bar{r}_t\theta}{\sigma^2 k} - \frac{1}{2} \frac{(\theta + \bar{\lambda})^2}{\sigma^2 k} - \frac{1}{2} \frac{(\theta - \bar{\lambda})^2}{\sigma^2 k} + O(k^{-\frac{3}{2}}) \\ & \geq \frac{\bar{\lambda}}{\sigma\sqrt{k}} + \frac{2\bar{r}_t^2}{\sigma^2 k} - \frac{1}{2} \frac{(\bar{r}_t + \bar{\lambda})^2}{\sigma^2 k} - \frac{1}{2} \frac{(\bar{r}_t - \bar{\lambda})^2}{\sigma^2 k} + O(k^{-\frac{3}{2}}) - \frac{2}{tk} \log(\alpha) \end{aligned} \quad (10)$$

All the terms involving $\bar{\lambda}$ cancel from this inequality. Multiplying by $\sigma^2 k$ thus yields

$$(\bar{r}_t - \theta)^2 \leq \sigma^2 \frac{-2 \log \alpha}{t} + O(k^{-\frac{1}{2}}), \quad (11)$$

which implies (9). ■

Finally, consider the set of one-step-ahead beliefs $\mathcal{P}_{t,k}(\bar{r}_t)$. Following (4), it contains all likelihoods of the type (5) for some $\theta \in \mathcal{M}_{t,k}^\alpha(s^t)$ and $|\lambda| < \bar{\lambda}$. It can thus

be summarized by an interval of probabilities for the high state next period, denoted $\left[\underline{p}_k(\bar{r}_t), \bar{p}_k(\bar{r}_t)\right]$, with both bounds strictly increasing in the sample mean \bar{r}_t :

$$\begin{aligned}\underline{p}_k(\bar{r}_t) &= \frac{1}{2} + \frac{1}{2} \frac{\underline{\theta}_k(\bar{r}_t) - \bar{\lambda}}{\sigma\sqrt{k}} \\ \bar{p}_k(\bar{r}_t) &= \frac{1}{2} + \frac{1}{2} \frac{\bar{\theta}_k(\bar{r}_t) + \bar{\lambda}}{\sigma\sqrt{k}}.\end{aligned}\tag{12}$$

To get an idea about the shrinkage of the interval of possible equity premia, consider the lowest and highest mean log returns per month. In the limit as $k \rightarrow \infty$,

$$\left[\lim_{k \rightarrow \infty} (\underline{\theta}_k(\bar{r}_t) - \bar{\lambda}), \lim_{k \rightarrow \infty} (\bar{\theta}_k(\bar{r}_t) + \bar{\lambda})\right] = [\bar{r}_t - \bar{\lambda} - t^{-\frac{1}{2}}\sigma b_\alpha, \bar{r}_t + \bar{\lambda} + t^{-\frac{1}{2}}\sigma b_\alpha].$$

3 Optimal portfolio weights

It is helpful to begin with the maximization step in (1), given some arbitrary probability $0 < p < 1$ for the high state. The optimal weight on stocks is

$$w(p) = \frac{e^{r^f/k} \left(p \left(e^{\sigma/\sqrt{k}} - e^{r^f/k} \right) + (1-p) \left(e^{-\sigma/\sqrt{k}} - e^{r^f/k} \right) \right)}{\left(e^{\sigma/\sqrt{k}} - e^{r^f/k} \right) \left(e^{r^f/k} - e^{-\sigma/\sqrt{k}} \right)}\tag{13}$$

Taylor expansions of the exponential terms lead to

$$\begin{aligned}w(p) &= \left(1 + \frac{r^f}{k} + O\left(\frac{1}{k^2}\right) \right) \times \\ &\frac{p \left(\frac{\sigma}{\sqrt{k}} + \frac{1}{2} \frac{\sigma^2}{k} - \frac{r^f}{k} + O\left(\frac{1}{k^{\frac{3}{2}}}\right) \right) + (1-p) \left(-\frac{\sigma}{\sqrt{k}} + \frac{1}{2} \frac{\sigma^2}{k} - \frac{r^f}{k} + O\left(\frac{1}{k^{\frac{3}{2}}}\right) \right)}{\left(\frac{\sigma}{\sqrt{k}} + \frac{1}{2} \frac{\sigma^2}{k} - \frac{r^f}{k} + O\left(\frac{1}{k^{\frac{3}{2}}}\right) \right) \left(\frac{r^f}{k} + \frac{\sigma}{\sqrt{k}} - \frac{1}{2} \frac{\sigma^2}{k} + O\left(\frac{1}{k^{\frac{3}{2}}}\right) \right)}\end{aligned}\tag{14}$$

$$= \frac{(2p-1)\sigma/\sqrt{k} + \frac{1}{2}\sigma^2/k - r^f/k + O\left(k^{-\frac{3}{2}}\right)}{\sigma^2/k + O\left(k^{-\frac{3}{2}}\right)}.\tag{15}$$

We are now ready to derive optimal portfolio weights. Begin with the Bayesian case:

Proposition S3. *The optimal Bayesian portfolio weight for horizon $T > 0$ is*

$$\lim_{k \rightarrow \infty} \omega_{t,k,T}^*(\bar{r}_t) = \frac{\bar{r}_t + \frac{1}{2}\sigma^2 - r^f}{\sigma^2} =: \omega_t^{bay}.$$

Proof. Consider the optimal portfolio weight for the Bayesian investor with horizon $T > 0$. There is no minimization step in (1); as a result, the objective function on the

left hand side is the sum of (i) the objective for a myopic investor and (ii) continuation utility, which depends on the investment horizon T , but is independent of the optimal portfolio weight. The optimal weight is therefore independent of the horizon T . The optimal weight is now obtained by evaluating $w(p)$ at $p = \phi_t = \phi_k(\bar{r}_t)$:

$$\begin{aligned}\omega_{t,k,T}^*(\bar{r}_t) &= w(\phi_k(\bar{r}_t)) \\ &= \frac{(2\phi_k(\bar{r}_t) - 1)\sigma/\sqrt{k} + \frac{1}{2}\sigma^2/k - r^f/k + O(k^{-\frac{3}{2}})}{\sigma^2/k + O(k^{-\frac{3}{2}})} \\ &= \frac{\bar{r}_t/k + \frac{1}{2}\sigma^2/k - r^f/k + O(k^{-\frac{3}{2}})}{\sigma^2/k + O(k^{-\frac{3}{2}})},\end{aligned}$$

where the third equality uses the definition of ϕ_k (recall (2)). Taking the limit as $k \rightarrow \infty$ delivers the Bayesian solution. ■

Consider next the problem with a nondegenerate set of one-step-ahead conditionals, but assume $T = 1/k$. In (12) above, the belief set $\mathcal{P}_{t,k}(\bar{r}_t)$ was defined in terms of an interval $[\underline{p}_k(\bar{r}_t), \bar{p}_k(\bar{r}_t)]$ for the probability of the high state, with bounds strictly increasing in \bar{r}_t .

Proposition S4. *The optimal portfolio weight of the myopic ambiguity averse investor ($T = 1/k$) is*

$$\begin{aligned}\lim_{k \rightarrow \infty} \omega_{t,k,1/k}^*(\bar{r}_t) &= \sigma^{-2} \max \left\{ \bar{r}_t + \frac{1}{2}\sigma^2 - r^f - \left(\bar{\lambda} + t^{-\frac{1}{2}}\sigma b_\alpha \right), 0 \right\} \\ &\quad + \sigma^{-2} \min \left\{ \bar{r}_t + \frac{1}{2}\sigma^2 - r^f + \bar{\lambda} + t^{-\frac{1}{2}}\sigma b_\alpha, 0 \right\} \\ &= \max \left\{ \omega_t^{bay} - \sigma^{-2} \left(\bar{\lambda} + t^{-\frac{1}{2}}\sigma b_\alpha \right), 0 \right\} \\ &\quad + \min \left\{ \omega_t^{bay} + \sigma^{-2} \left(\bar{\lambda} + t^{-\frac{1}{2}}\sigma b_\alpha \right), 0 \right\}.\end{aligned}$$

Proof. For given p , the optimal weight is $w(p)$. To solve the minimization step, substituting $w(p)$ back into the objective (1). We now need to find $p \in [\underline{p}_k(\bar{r}_t), \bar{p}_k(\bar{r}_t)]$ to minimize

$$\begin{aligned}g(p) &= p \log \left(e^{r^f/k} + w(p) \left(e^{\sigma/\sqrt{k}} - e^{r^f/k} \right) \right) \\ &\quad + (1-p) \log \left(e^{r^f/k} + w(p) \left(e^{-\sigma/\sqrt{k}} - e^{r^f/k} \right) \right) \\ &= r^f/k + p \log p + (1-p) \log(1-p) + \log \left(e^{\sigma/\sqrt{k}} - e^{-\sigma/\sqrt{k}} \right) \\ &\quad - p \log \left(e^{r^f/k} - e^{-\sigma/\sqrt{k}} \right) - (1-p) \log \left(e^{\sigma/\sqrt{k}} - e^{r^f/k} \right).\end{aligned}$$

The function g is strictly convex on $(0, 1)$ and achieves a minimum at

$$\hat{p}_k = \frac{e^{r^f/k} - e^{-\sigma/\sqrt{k}}}{e^{\sigma/\sqrt{k}} - e^{-\sigma/\sqrt{k}}}. \quad (16)$$

The minimizer \hat{p}_k is in the unit interval because we have assumed that $r^f < \sigma$. It is precisely the probability at which the one-step-ahead conditional equity premium $E[R(s_{t+1/k})] - R^f$ is equal to zero. The conditional premium appears also as the bracketed term in the numerator of (13), so that $w(\hat{p}_k) = 0$. It follows that the solution to the minimization step is

$$p_k^*(\bar{r}_t) = \begin{cases} \bar{p}_k(\bar{r}_t) & \text{if } \hat{p}_k > \bar{p}_k(\bar{r}_t) \\ \hat{p}_k & \text{if } \hat{p}_k \in [\underline{p}_k(\bar{r}_t), \bar{p}_k(\bar{r}_t)] \\ \underline{p}_k(\bar{r}_t) & \text{if } \hat{p}_k < \underline{p}_k(\bar{r}_t). \end{cases}$$

Substituting into (13) and using $w(\hat{p}_k) = 0$, we can express the optimal portfolio weight as a function of the sample mean:

$$\omega_{t,k,1/k}^*(\bar{r}_t) = \begin{cases} w(\bar{p}_k(\bar{r}_t)) & \text{if } \hat{p}_k > \bar{p}_k(\bar{r}_t) \\ 0 & \text{if } \hat{p}_k \in [\underline{p}_k(\bar{r}_t), \bar{p}_k(\bar{r}_t)] \\ w(\underline{p}_k(\bar{r}_t)) & \text{if } \hat{p}_k < \underline{p}_k(\bar{r}_t). \end{cases}$$

We now compute $\lim_{k \rightarrow \infty} \omega_{t,k,1/k}^*(\bar{r})$. Since the functions \underline{p}_k and \bar{p}_k are strictly increasing, the nonparticipation region of the state space can be represented by an interval of sample means $[\bar{r}_{lo}(k), \bar{r}_{up}(k)]$. If $\bar{r}_t > \bar{r}_{up}(k)$ the evidence about the equity premium is so positive that investment in stocks is positive even under the lowest probability for the high state. The upper bound $\bar{r}_{up}(k)$ is the unique solution to $\underline{p}_k(\bar{r}_t) = \hat{p}_k$ or, equivalently,

$$\left(2\underline{p}_k(\bar{r}_t) - 1\right) \sigma/\sqrt{k} = (2\hat{p}_k - 1) \sigma/\sqrt{k} \quad (17)$$

Use (12) and 11) above to rewrite the left hand side as:

$$\begin{aligned} \left(2\underline{p}_k(\bar{r}_t) - 1\right) \sigma/\sqrt{k} &= \frac{1}{k} \left(\bar{r}_t - \bar{\lambda} - \sqrt{\sigma^2 \frac{-2 \log \alpha}{t} + O(k^{-\frac{1}{2}})} \right) \\ &= \frac{1}{k} \left(\bar{r}_t - \sigma b_\alpha/\sqrt{t} - \bar{\lambda} \right) + O(k^{-\frac{5}{4}}). \end{aligned} \quad (18)$$

Substitute (18) and (16) into (17) to obtain

$$\frac{1}{k} \left(\bar{r}_t - \sigma b_\alpha/\sqrt{t} - \bar{\lambda} \right) + O(k^{-\frac{5}{4}}) = \frac{2e^{r^f/k} - e^{\sigma/\sqrt{k}} - e^{-\sigma/\sqrt{k}}}{e^{\sigma/\sqrt{k}} - e^{-\sigma/\sqrt{k}}} \frac{\sigma}{\sqrt{k}}.$$

Taylor expansions of the exponential terms on the right side around 1 lead to

$$\frac{1}{k} \left(\bar{r}_t - \sigma b_\alpha / \sqrt{t} - \bar{\lambda} \right) + O \left(k^{-\frac{5}{4}} \right) = \frac{2r^f/k - \sigma^2/k}{2\sigma/\sqrt{k}} \frac{\sigma}{\sqrt{k}},$$

so that the upper bound of the nonparticipation region can be written as

$$\bar{r}_{up}(k) = \sigma b_\alpha / \sqrt{t} + \bar{\lambda} + r^f - \frac{1}{2} \sigma^2 + O \left(k^{-\frac{1}{4}} \right). \quad (19)$$

The lower bound $\bar{r}_{lo}(k)$ is the unique solution to $\bar{p}_k(\bar{r}_t) = \hat{p}_k$. By an argument similar to the one above, it can be written as

$$\bar{r}_{lo}(k) = -\sigma b_\alpha / \sqrt{t} - \bar{\lambda} + r^f - \frac{1}{2} \sigma^2 + O \left(k^{-\frac{1}{4}} \right). \quad (20)$$

Consider next the optimal weight in the case $\underline{p}_k(\bar{r}_t) > \hat{p}_k$. Evaluating $w(p)$ from (15) at $p = \underline{p}_k(\bar{r}_t)$ and replacing the first term in the numerator using (18) yields

$$w \left(\underline{p}_k(\bar{r}_t) \right) = \frac{\bar{r}_t - \sigma b_\alpha / \sqrt{t} - \bar{\lambda} + \frac{1}{2} \sigma^2 - r^f + O \left(k^{-\frac{1}{4}} \right)}{\sigma^2 + O \left(k^{-\frac{1}{2}} \right)} = \frac{\bar{r}_t - \bar{r}_{up}(k) + O \left(k^{-\frac{1}{4}} \right)}{\sigma^2 + O \left(k^{-\frac{1}{2}} \right)}. \quad (21)$$

The case $\hat{p}_k > \bar{p}_k(\bar{r}_t)$ is analogous, but with $\bar{r}_{up}(k)$ replaced by $\bar{r}_{lo}(k)$.

Combine the bounds (19)-(20) and the formulas for the optimal weight to deduce that the sequence of functions $\left\{ \omega_{t,k,1/k}^*(\bar{r}_t) \right\}$ converges pointwise to

$$\lim_{k \rightarrow \infty} \omega_{t,k,1/k}^*(\bar{r}) = \begin{cases} \frac{\bar{r}_t + \sigma b_\alpha / \sqrt{t} + \bar{\lambda} \frac{1}{2} \sigma^2 - r}{\sigma^2} & \text{if } \bar{r}_t < r^f - \frac{1}{2} \sigma^2 - \sigma b_\alpha / \sqrt{t} - \bar{\lambda} \\ 0 & \text{if } |\bar{r}_t + \frac{1}{2} \sigma^2 - r^f| < \sigma b_\alpha / \sqrt{t} + \bar{\lambda} \\ \frac{\bar{r}_t - \sigma b_\alpha / \sqrt{t} - \bar{\lambda} + \frac{1}{2} \sigma^2 - r}{\sigma^2} & \text{if } \bar{r}_t > r^f - \frac{1}{2} \sigma^2 + \sigma b_\alpha / \sqrt{t} + \bar{\lambda}. \end{cases}$$

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