

# MCMC

Numerical methods for Bayes

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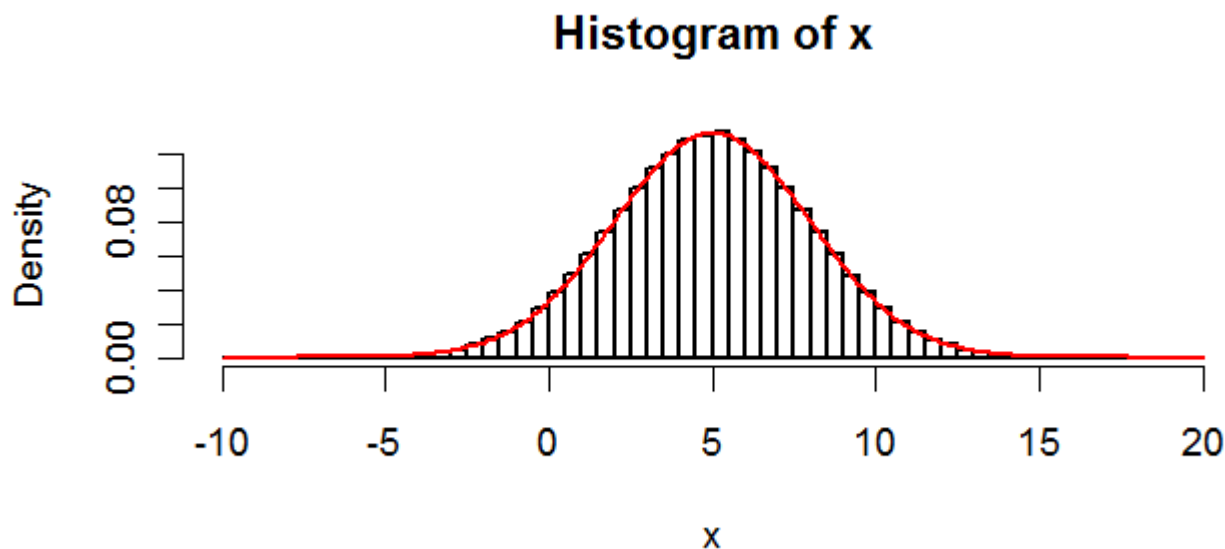
$$P(\theta|y) = \frac{P(y|\theta)P(\theta)}{\int_{-\infty}^{\infty} P(y|\theta)P(\theta)d\theta}$$

- Not just optimization
- Need to integrate denominator
  - Numerical Integration
- Would also like to know the mean, median, mode, variance, quantiles, confidence intervals, etc.

# Idea:

## Random samples from the posterior

- Approximate PDF with the histogram
- Performs *Monte Carlo Integration*
- Allows all quantities of interest to be calculated from the sample (mean, quantiles, var, etc)



	TRUE	Sample
mean	5.000	5.000
median	5.000	5.004
var	9.000	9.006
Lower CI	-0.880	-0.881
Upper CI	10.880	10.872

# Outline

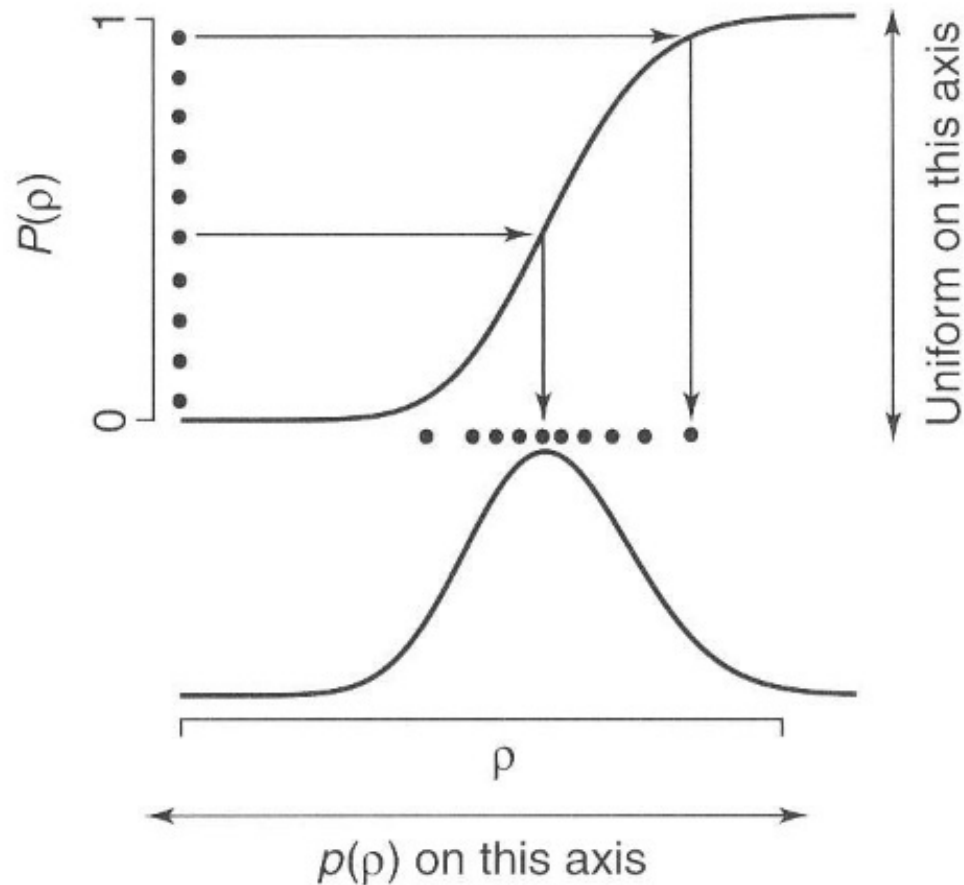
- **Different numerical techniques for sampling from the posterior**
  - **Inverse Distribution Sampling**
  - **Rejection Sampling & SMC**
  - **Markov Chain-Monte Carlo (MCMC)**
    - Metropolis
    - Metropolis-Hastings
    - Gibbs sampling
- Sampling conditionals vs full model
- Flexibility to specify complex models

# How do we generate a random number from a PDF?

- Exist for most standard distributions
- Posteriors often non-standard
- Indirect Methods
  - First sample from a different distribution
  - Rejection sampling, Metropolis, M-H
- Direct Methods
  - Inverse CDF
  - Univariate sampling of multivariate or conditional

# Inverse CDF sampling

- 1) Sample from a uniform distribution
- 2) Transform sample using inverse of CDF,  $F^{-1}(x)$



# Example: Exponential

- The exponential CDF is:  $F(x) = 1 - e^{-\lambda x}$
- We solve for  $F^{-1}$  as

$$p = 1 - e^{-\lambda x}$$

$$1 - p = e^{-\lambda x}$$

$$\ln(1 - p) = -\lambda x$$

$$x = F^{-1}(p) = -\frac{\ln(1 - p)}{\lambda}$$

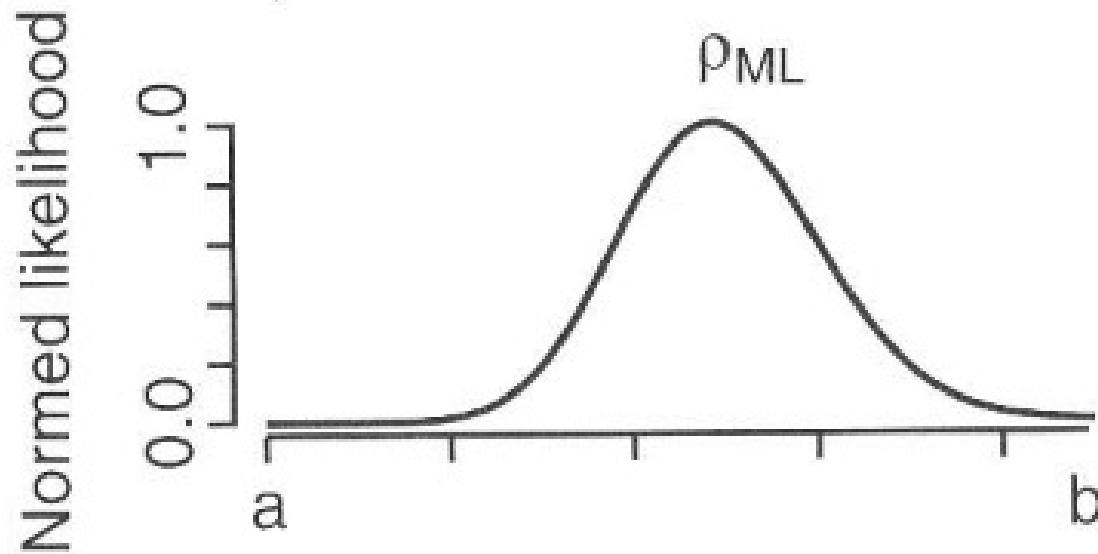
- Draw  $p \sim \text{Unif}(0, 1)$ , calculate  $x$

# Approximate inverse sampling

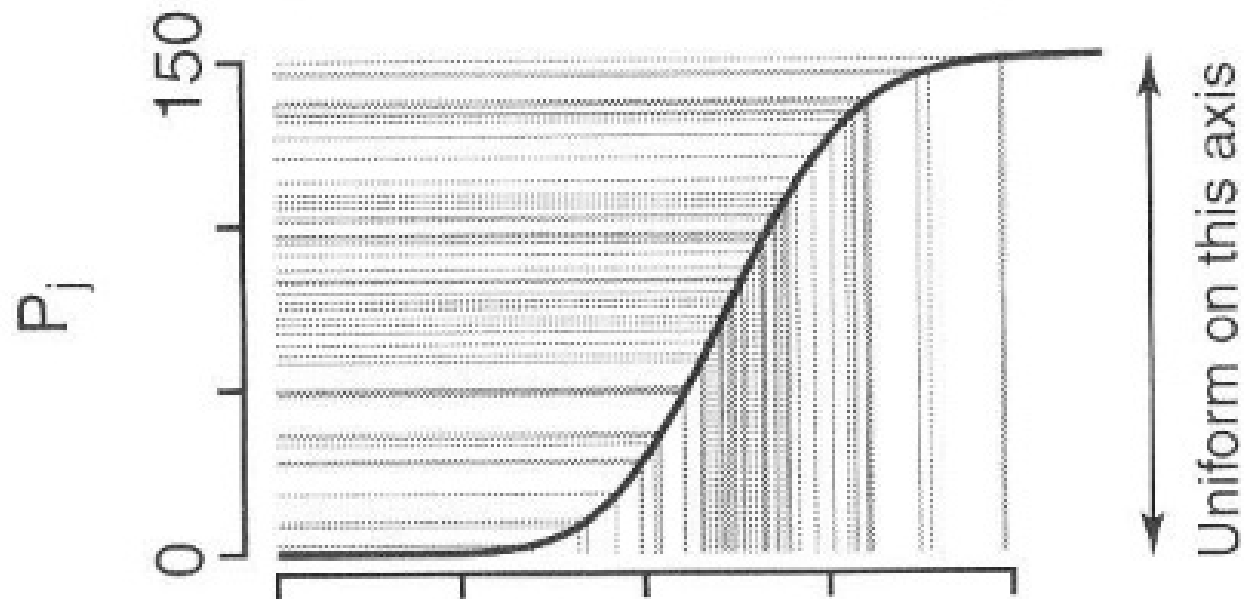
- Exact inverse sampling requires CDF & ability to solve for inverse
- Approximation
  - Solve for  $f(x)$  across a discrete sequence of  $x$
  - Determine cumulative sum to approx  $F(x)$
  - Draw  $Z \sim \text{unif}(0, \max)$
  - Find the value of  $x$  for which  $Z \approx \text{cumsum}(f(x))$
- Approximation performs integration as a Riemann sum



a) Likelihood normalized to maximum value



b) Cumulative likelihood ( $m = 100$ )



# Univariate sampling of multivariate or conditional distribution

- Multivariate
  - Multivariate normal based on Normal
  - Multinomial based on Binomial
- Conditional
  - Sample from the first distribution
  - Sample from the second conditioned on the first
  - Examples
    - $\text{NBin} = \text{Pois}(y|\lambda)\text{Gamma}(\lambda | a, b)$
    - $\text{Students } t = \text{Normal}(x | \mu, \sigma^2) \text{IG}(\sigma^2|a, b)$

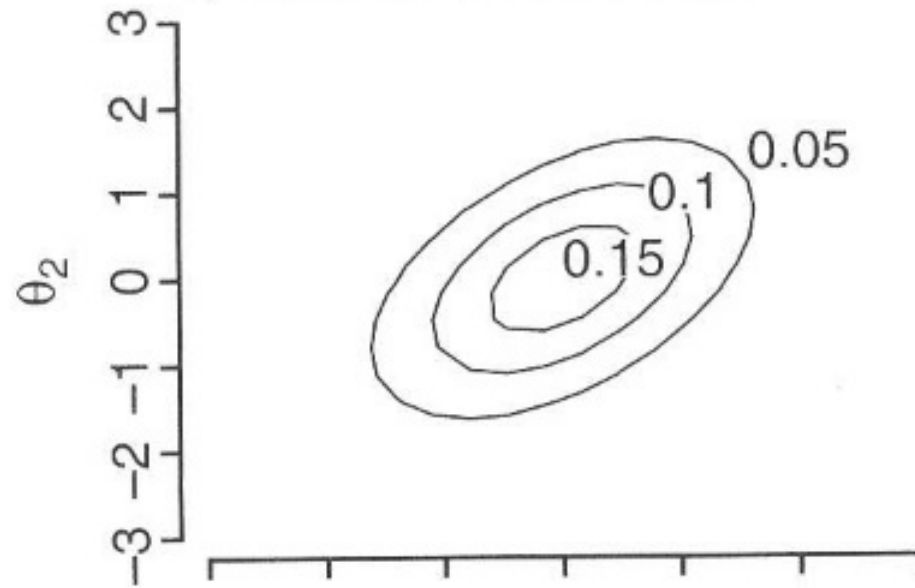
# Markov Chain Monte Carlo

- 1) Start from some initial parameter value
- 2) Evaluate the unnormalized posterior
- 3) Propose a new parameter value
- 4) Evaluate the new unnormalized posterior
- 5) Decide whether or not to accept the new value
- 6) Repeat 3-5

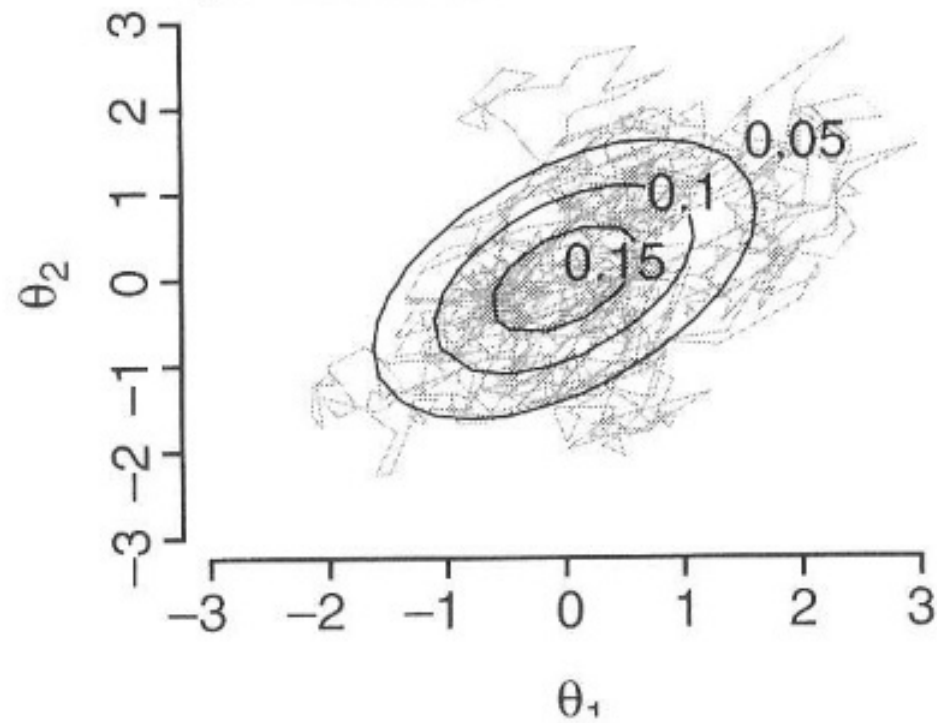
# Markov Chain Monte Carlo

- Looks remarkably similar to optimization
  - Evaluating posterior rather than just likelihood
  - “Repeat” does not have a stopping condition
  - **Criteria for accepting a proposed step**
    - Optimization – diverse variety of options but no “rule”
    - MCMC – stricter criteria for accepting
- Performs random walk through PDF
- Converges “in distribution” rather than to a single point

a) A bivariate normal distribution

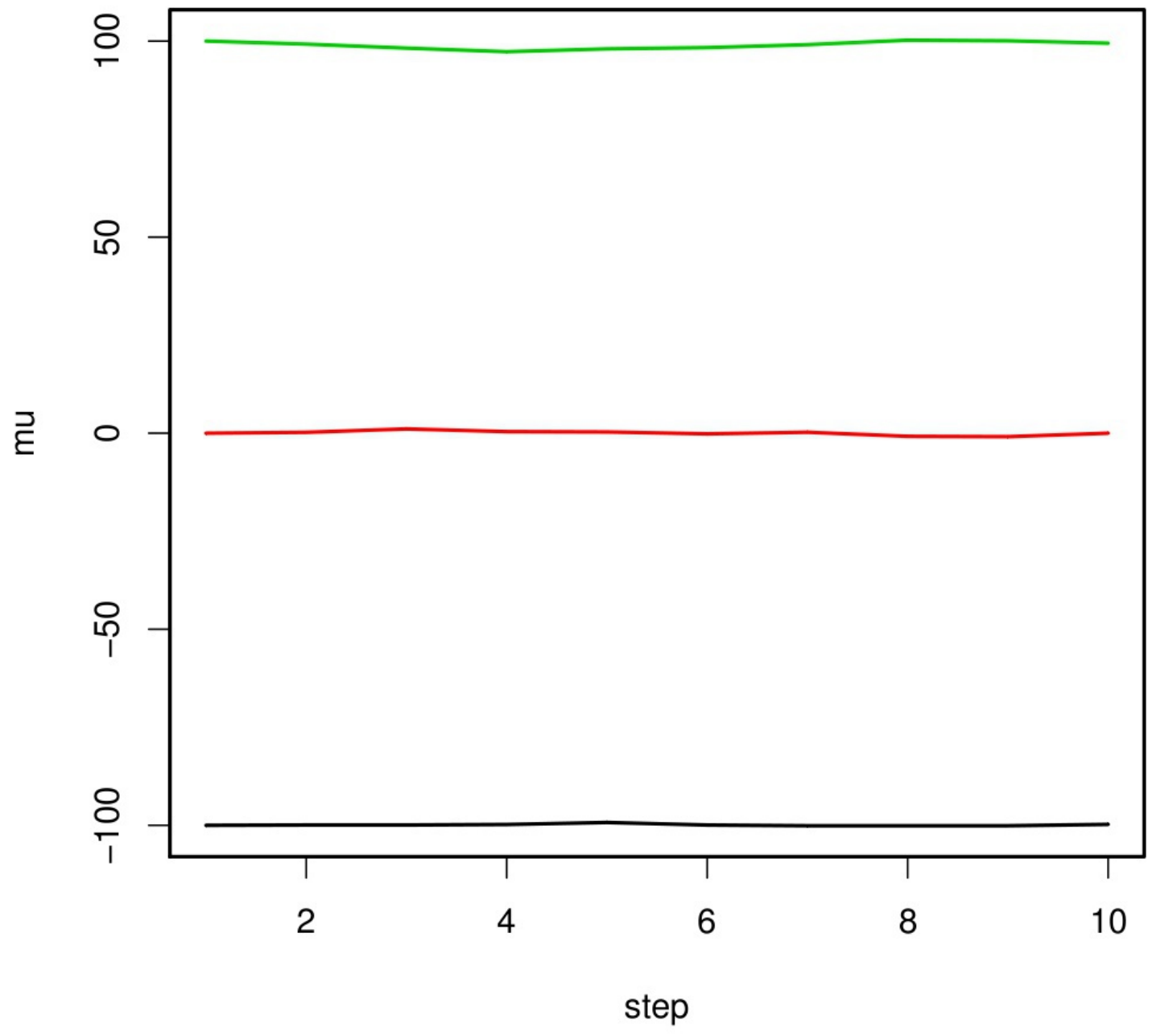


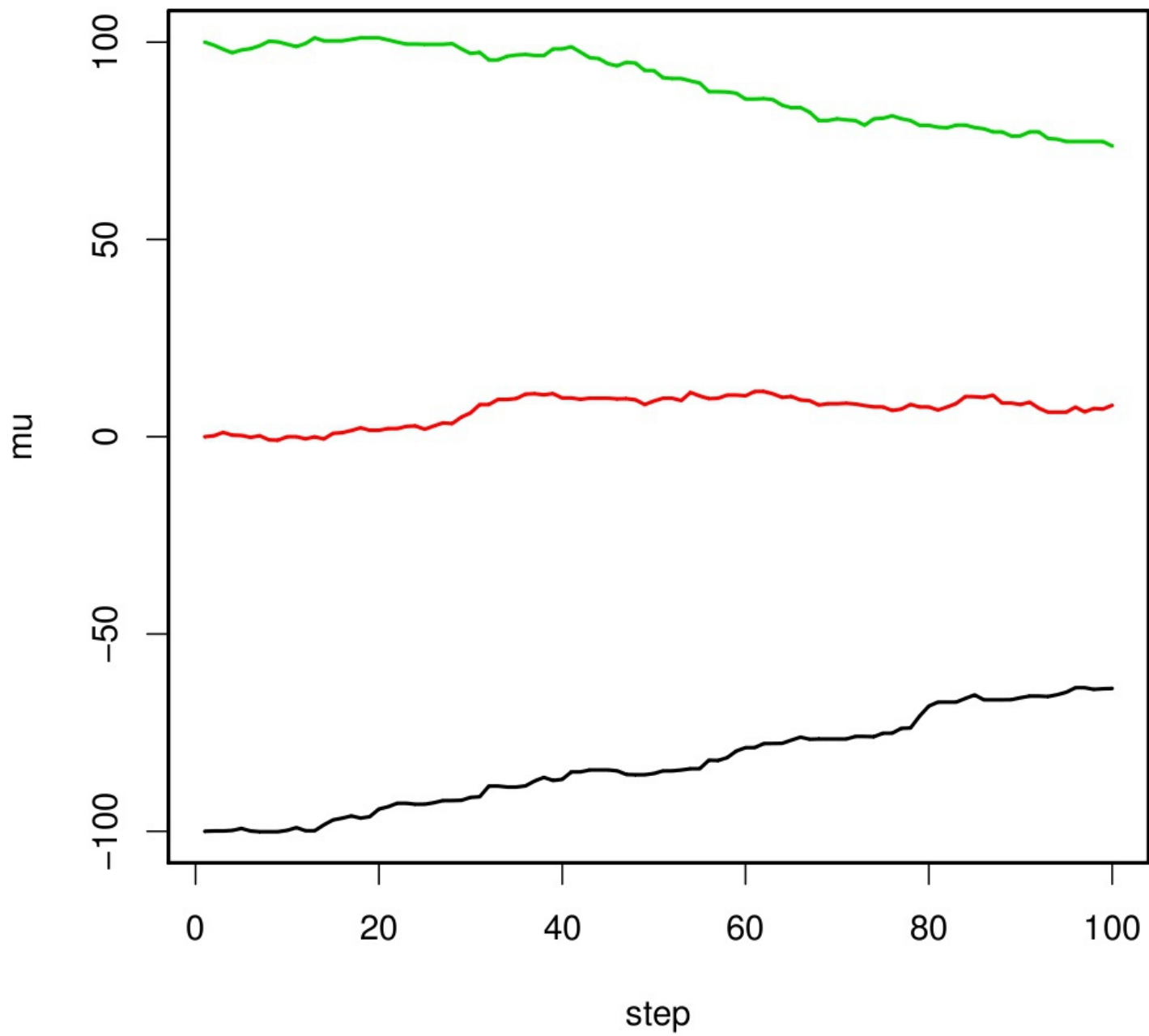
b) Simulation



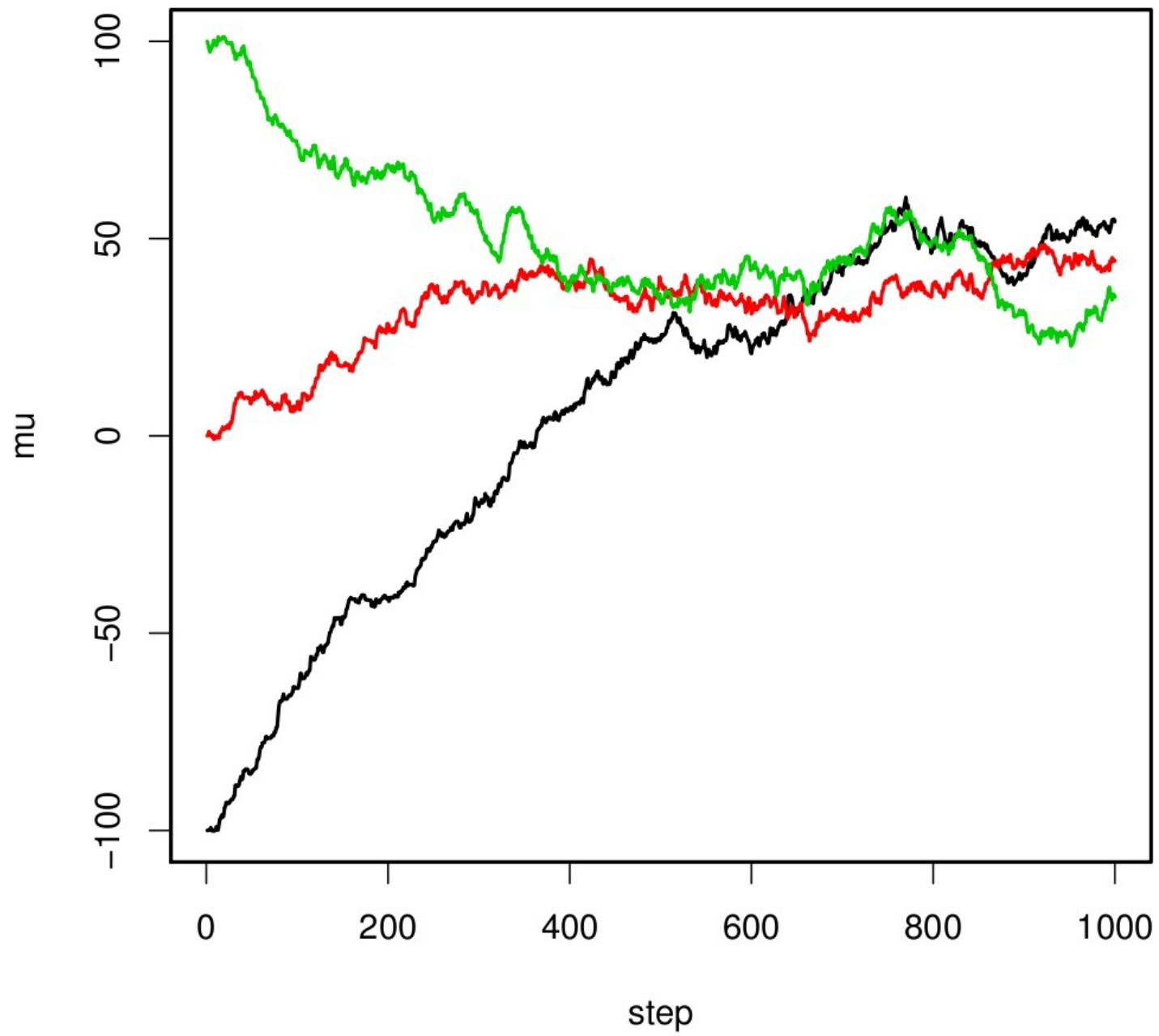
# Example

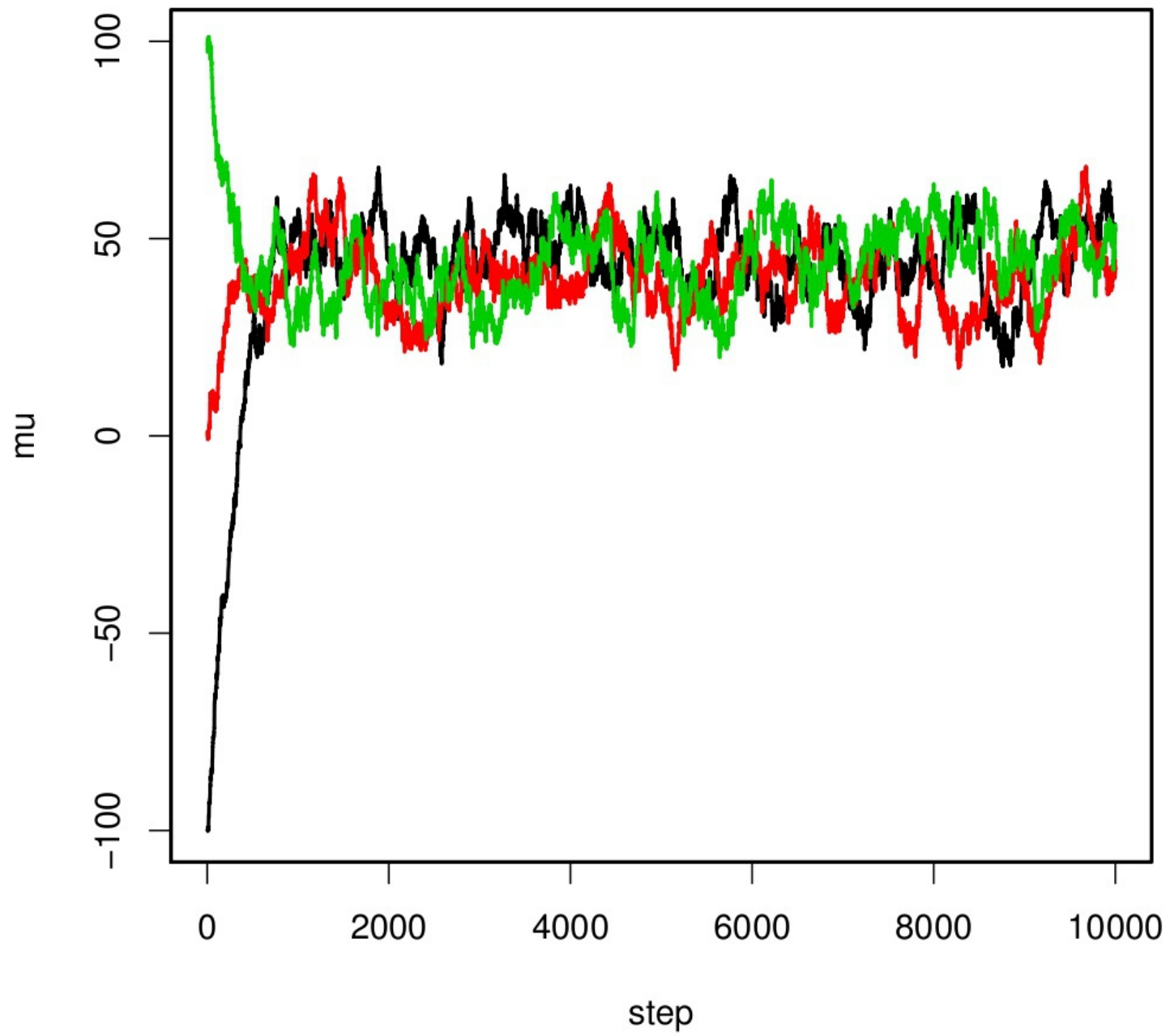
- Normal with known variance, unknown mean
  - Prior:  $N(53, 10000)$
  - Data:  $y = 43$
  - Known variance: 100
  - Initial conditions, 3 chains starting at -100, 0, 100



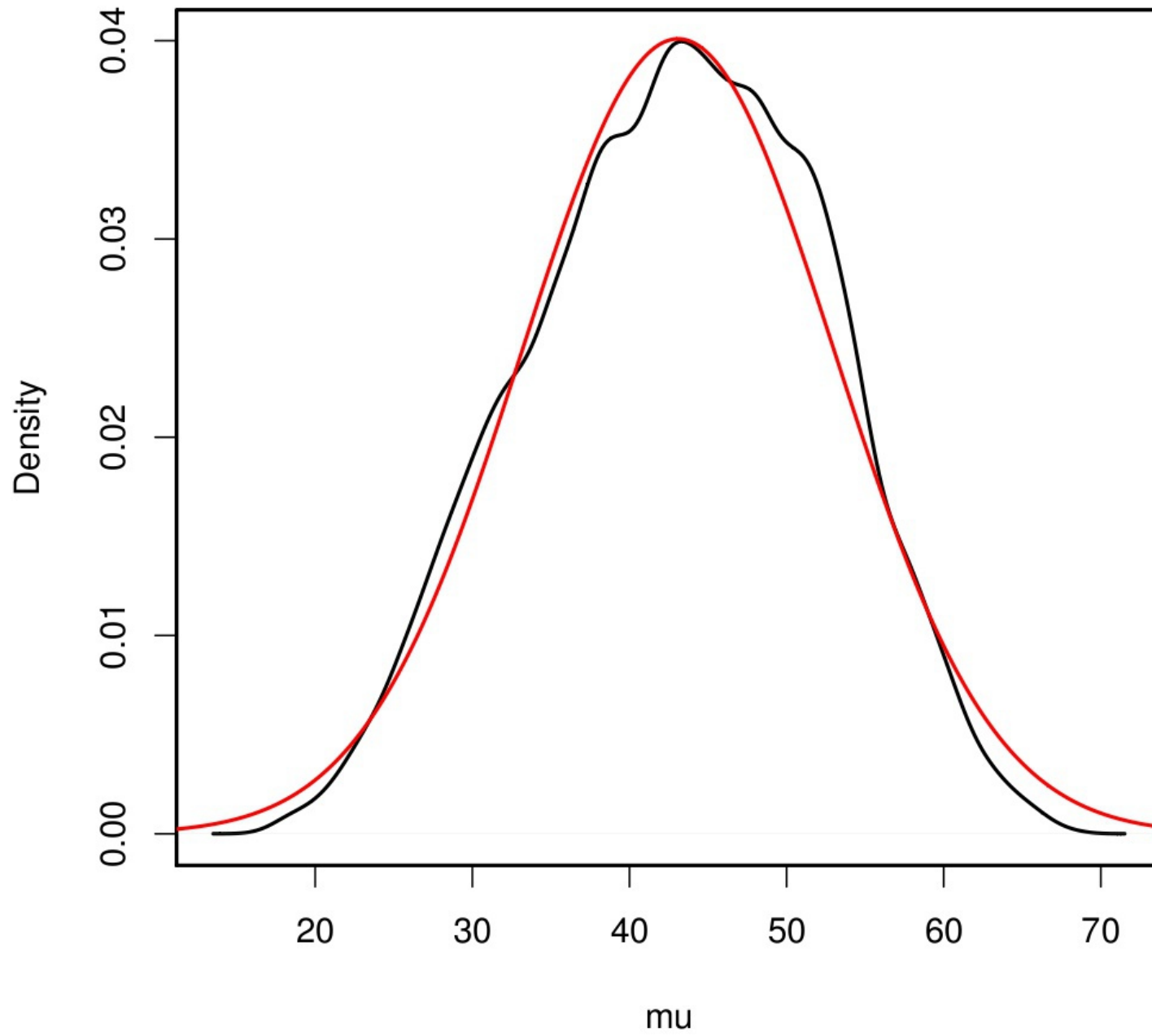








# MCMC Posterior Density



- Advantages

- Multi-dimensional
- Can be applied to
  - Whole joint PDF
  - Each dimension iteratively
  - Groups of parameters
- Simple
- Robust

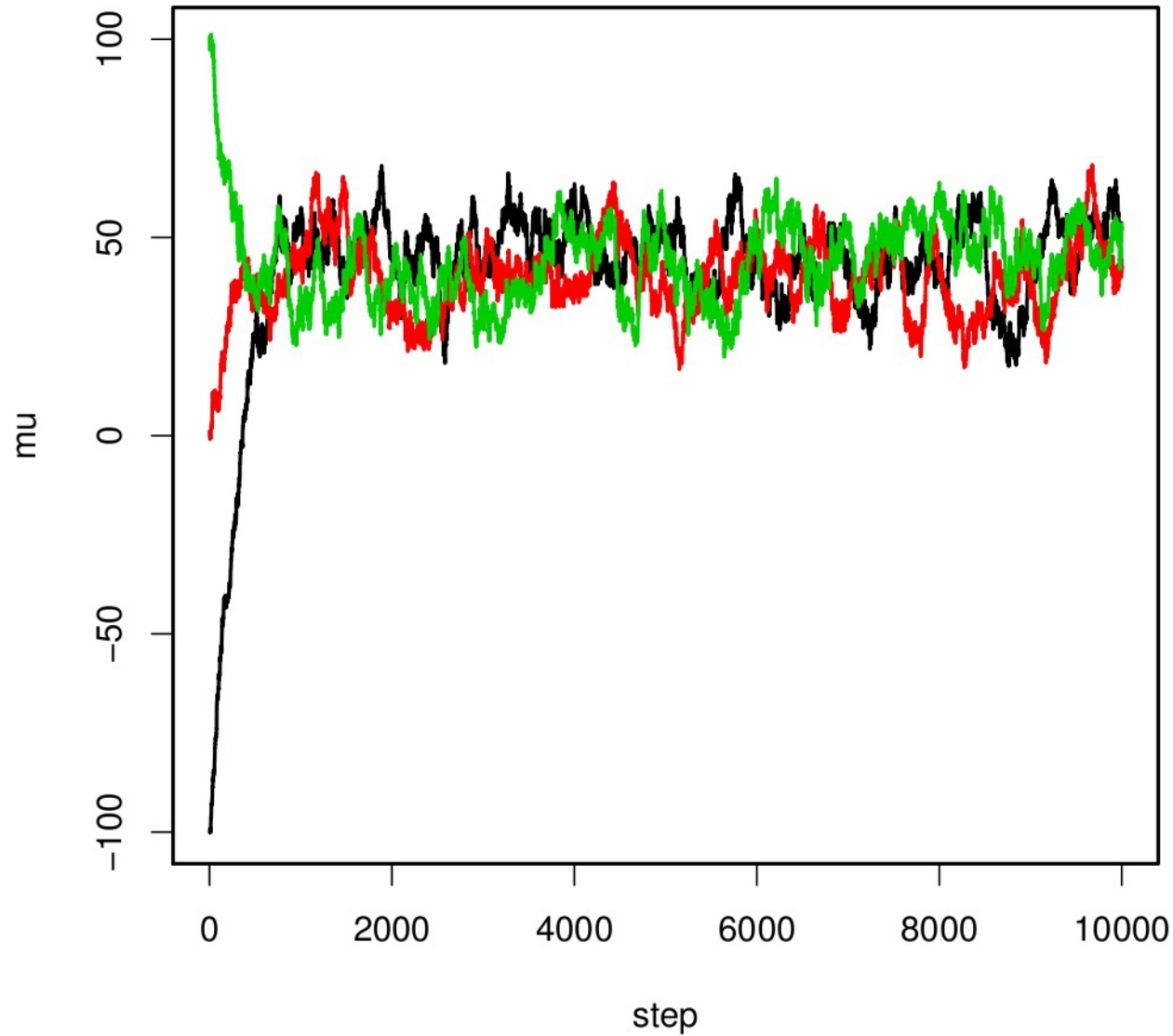
- Disadvantages

- Sequential samples not independent
- Computationally intensive
- Discard “Burn – in” period before convergence
- Assessing convergence

# Convergence

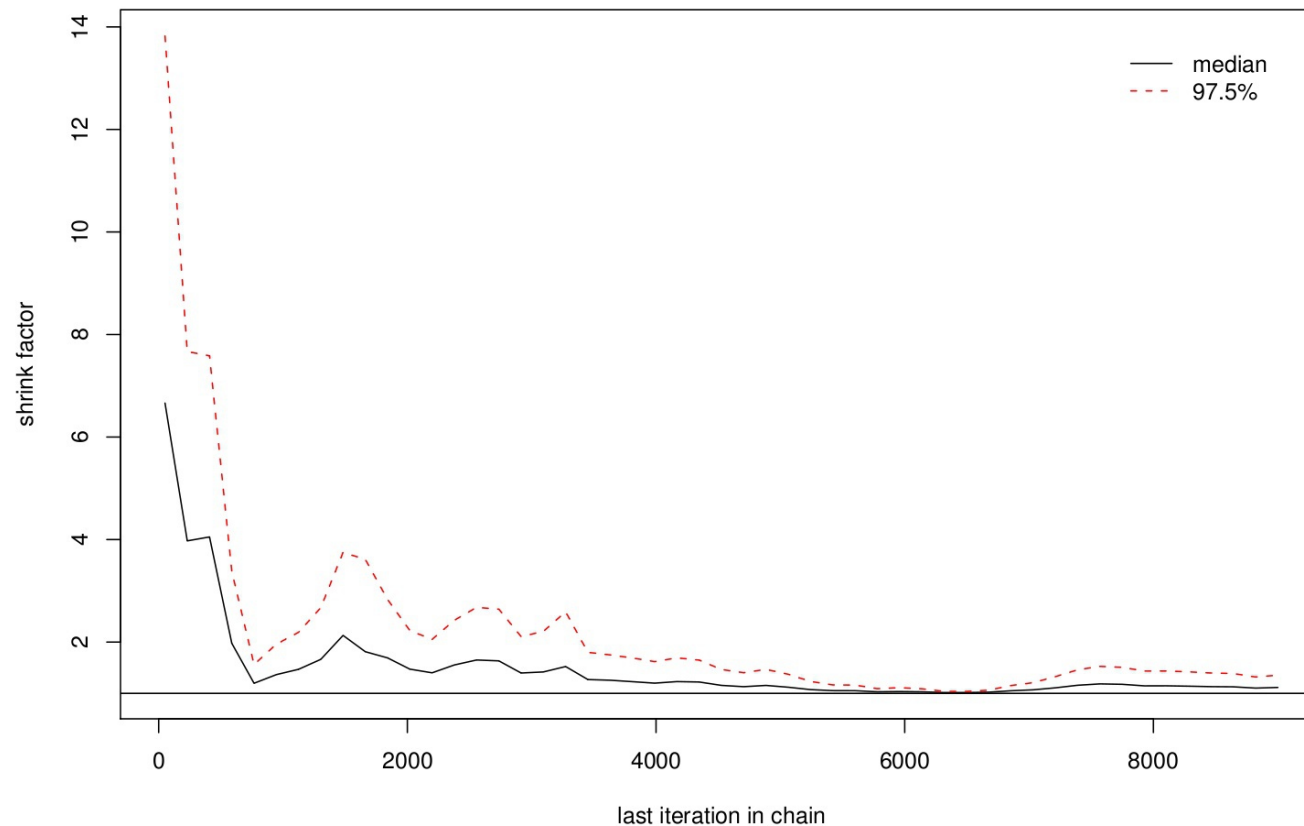
- Generally can not be “proved”
- Why MCMC can be “dangerous,” especially in the hands of the untrained
- Assessed by examining MCMC time-series
  - Visual inspection
  - Multiple chains
  - Convergence statistics
  - Acceptance rate
  - Auto-correlation

# Visual inspection / multiple chains



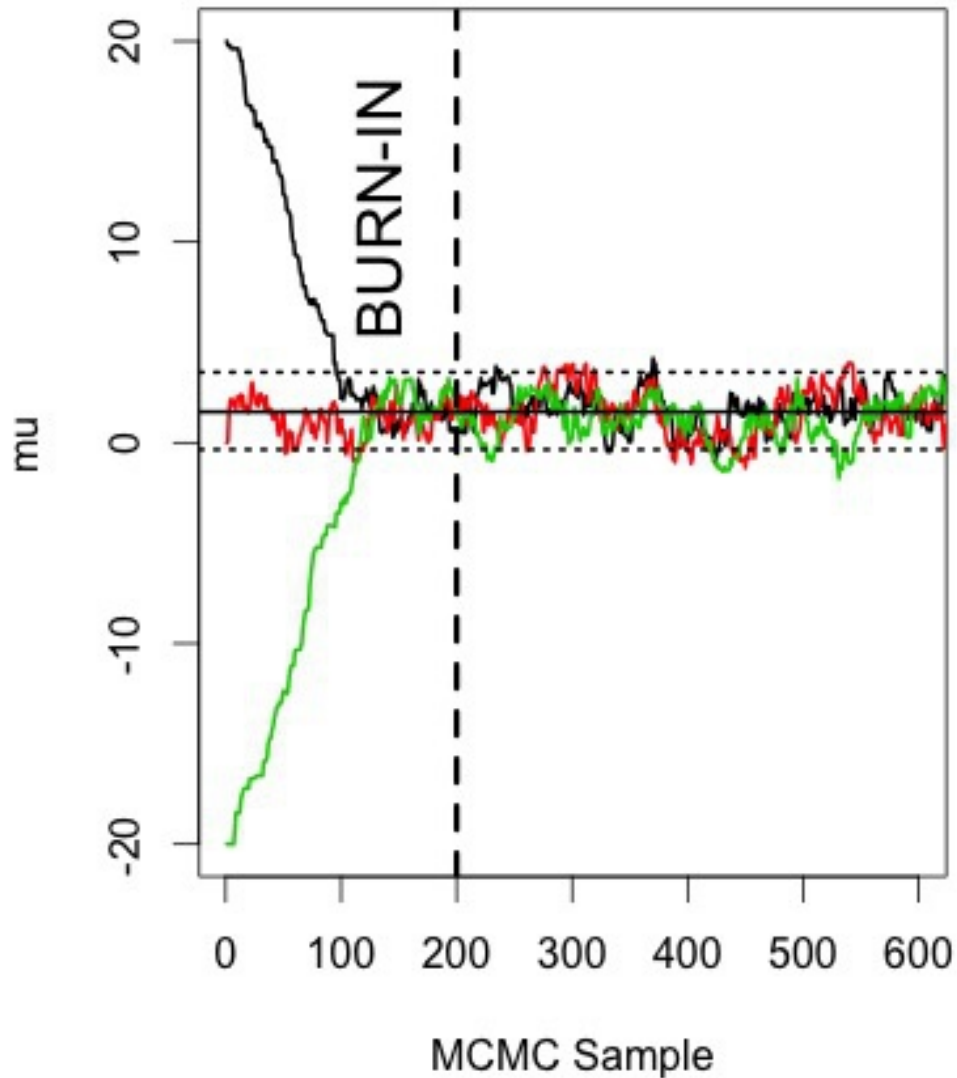
# Convergence Statistics

- Brooks Gelman Rubin
  - Within vs among chain variance
  - Should converge to 1

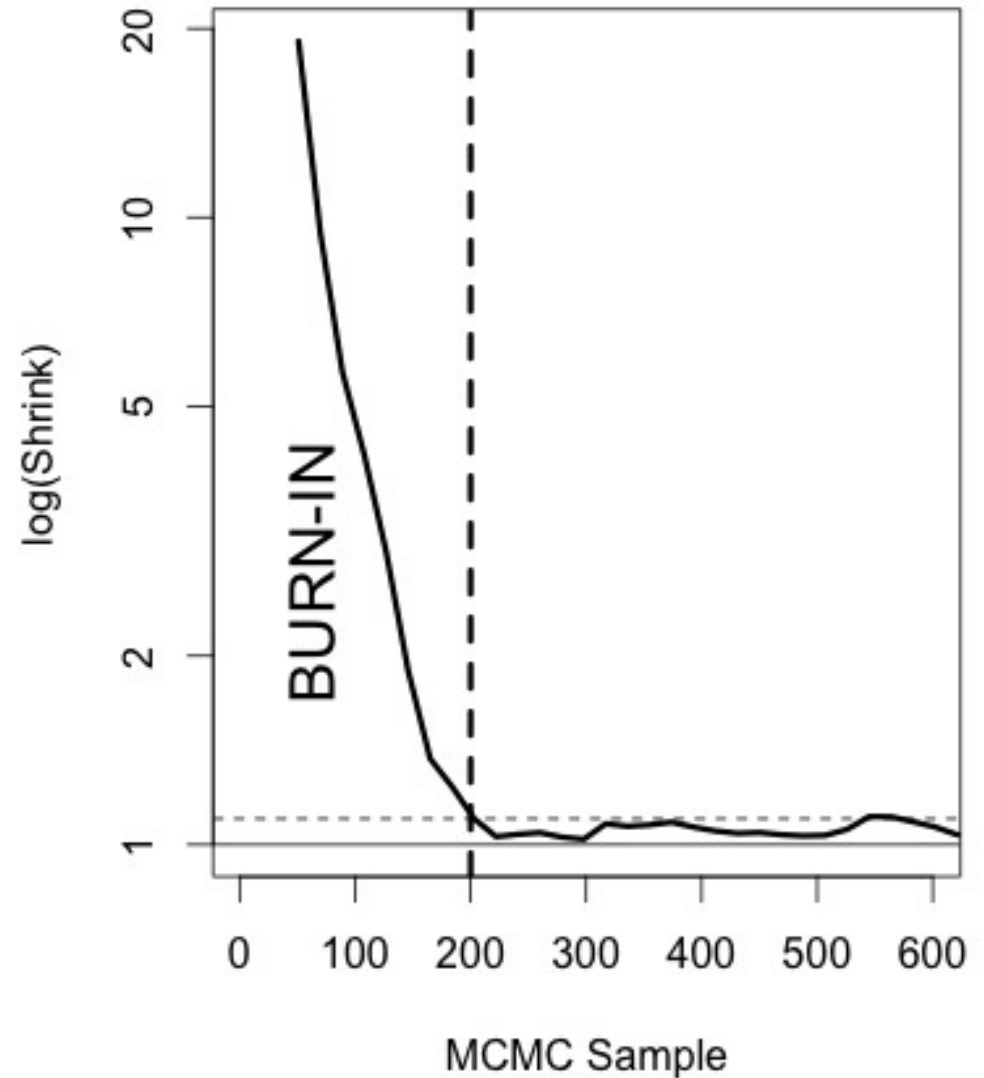


# Convergence Statistics

Trace Plot

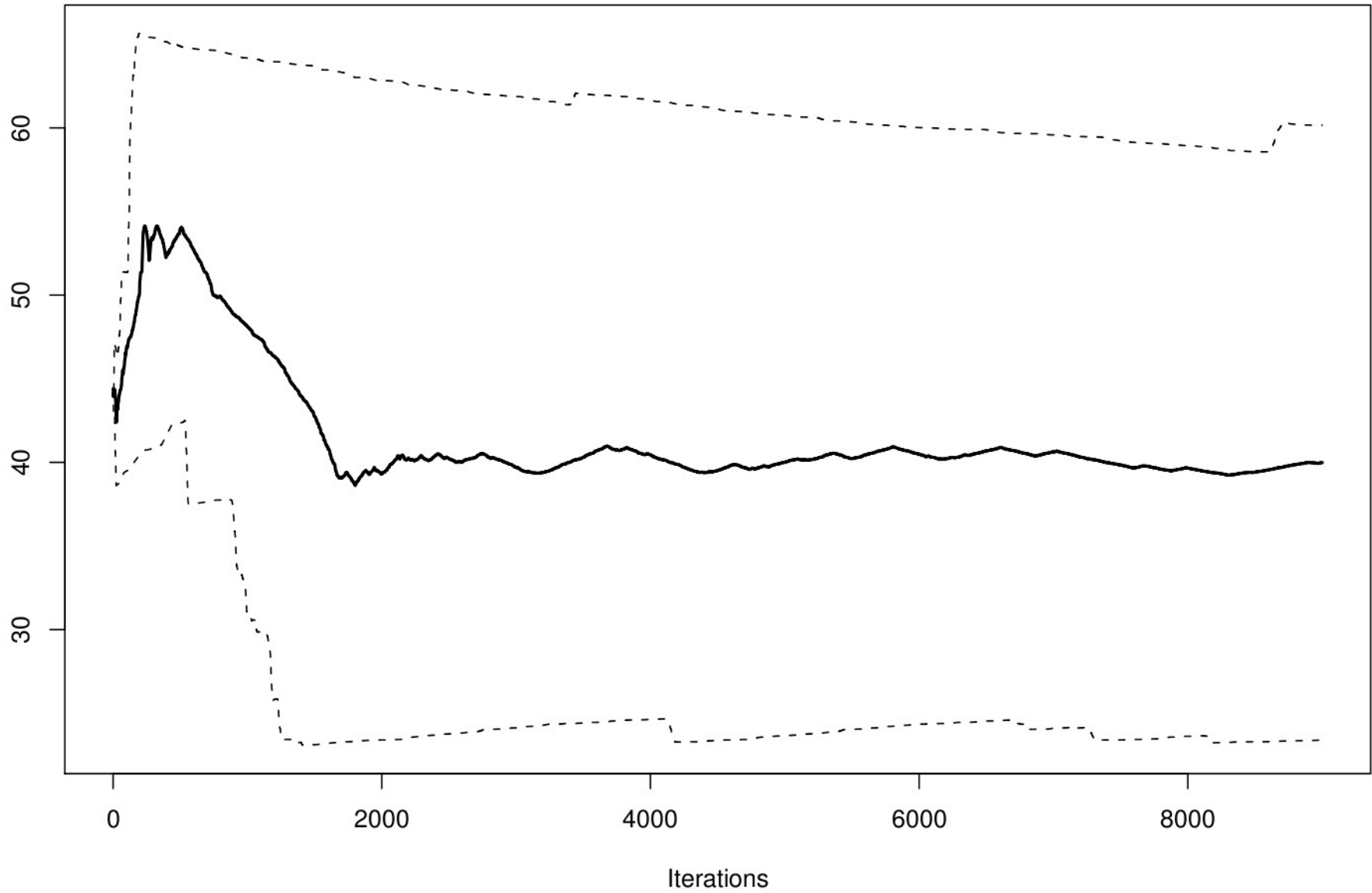


GBR Diagnostic

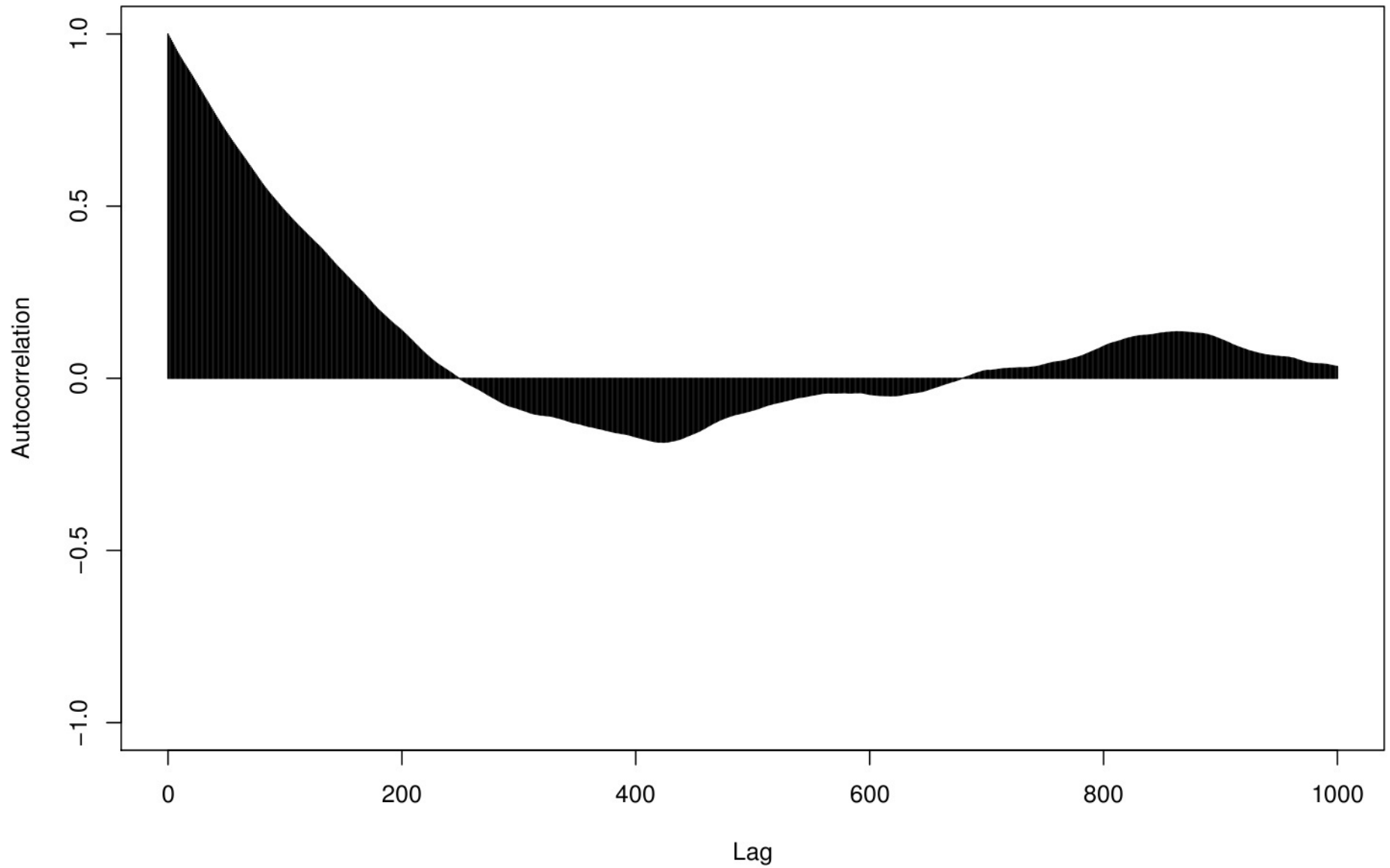




# Quantiles



# Autocorrelation



# Acceptance Rate

- Metropolis & Metropolis – Hastings
  - Aim for 30-70%
  - Too low = not mixing
  - Too high = small steps, slow mixing
  - Example: 97%
- Gibbs sampling
  - Always 100%

# Summary Statistics

## Analytical:

Mean	SD
43.09901	9.95037

## MCMC:

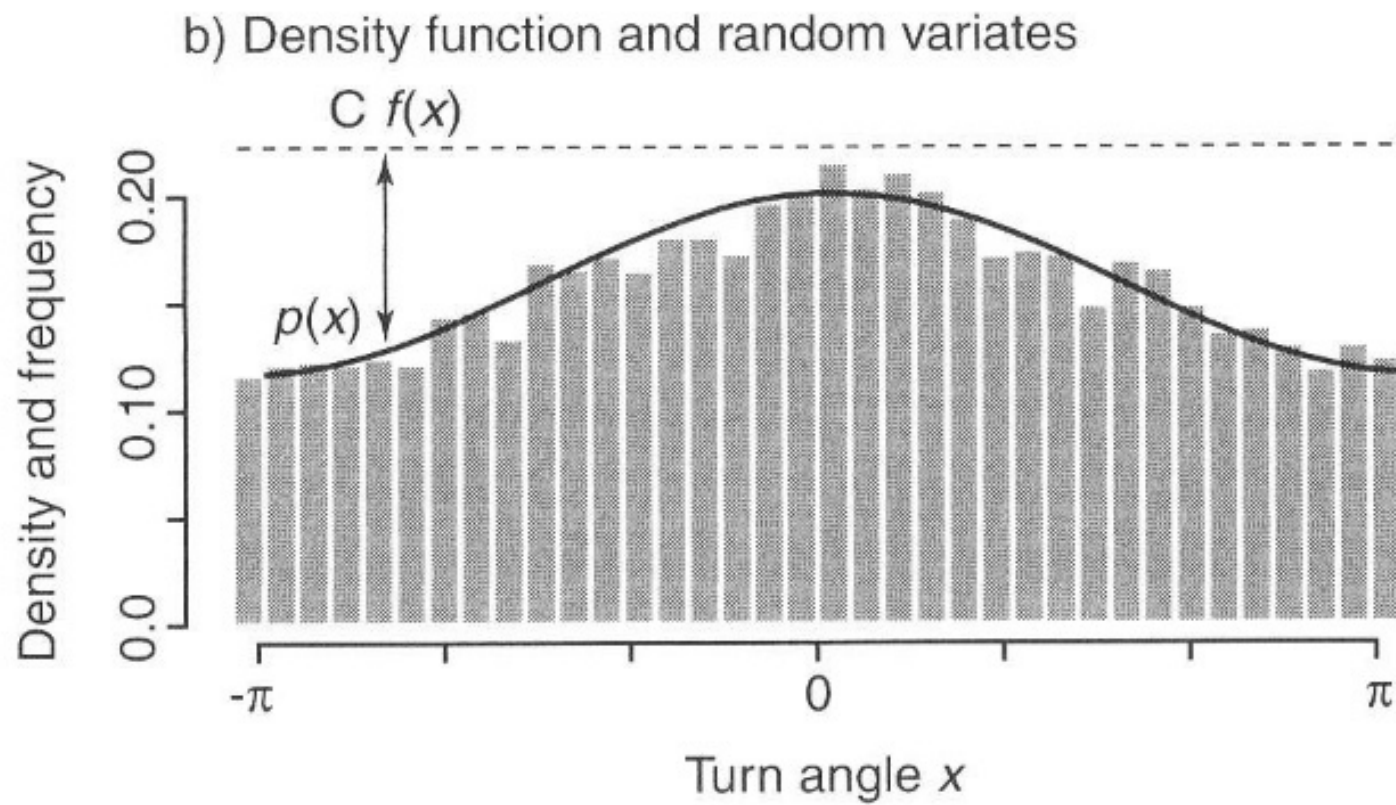
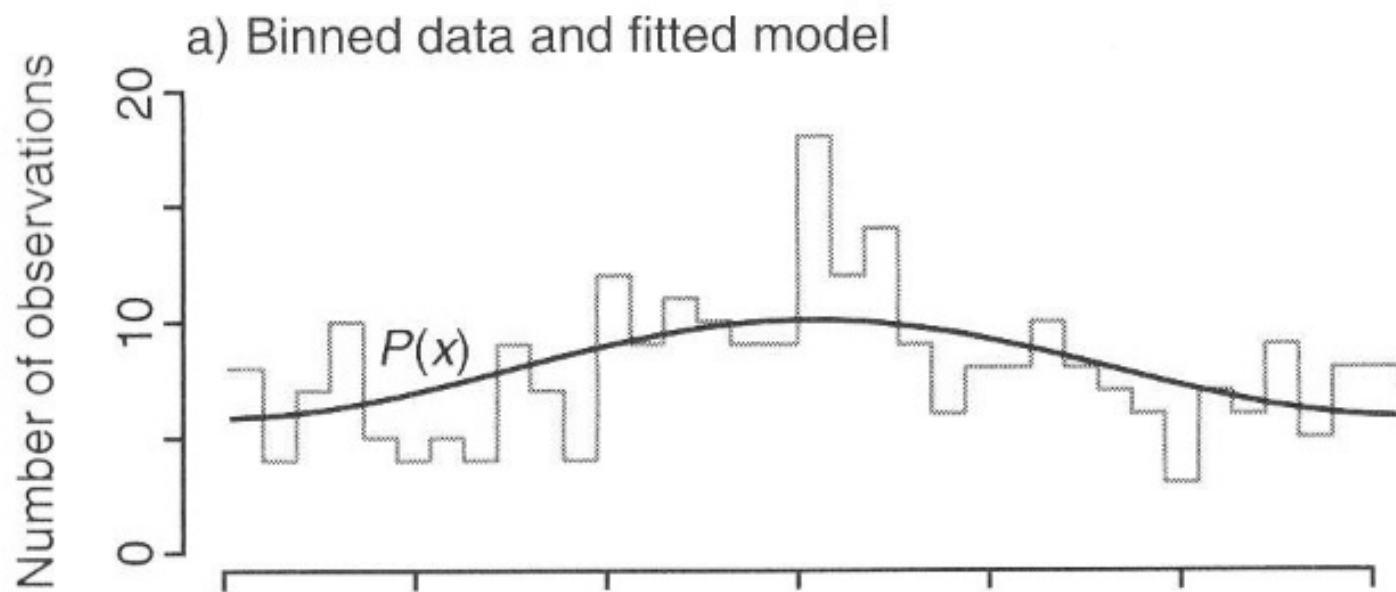
Mean	SD	Naive SE	Time-series
43.05504	9.28108	0.05648	0.74503

## Quantiles:

2.5%	25%	50%	75%	97.5%
24.98	36.46	43.39	49.99	60.01

# Rejection Sampling

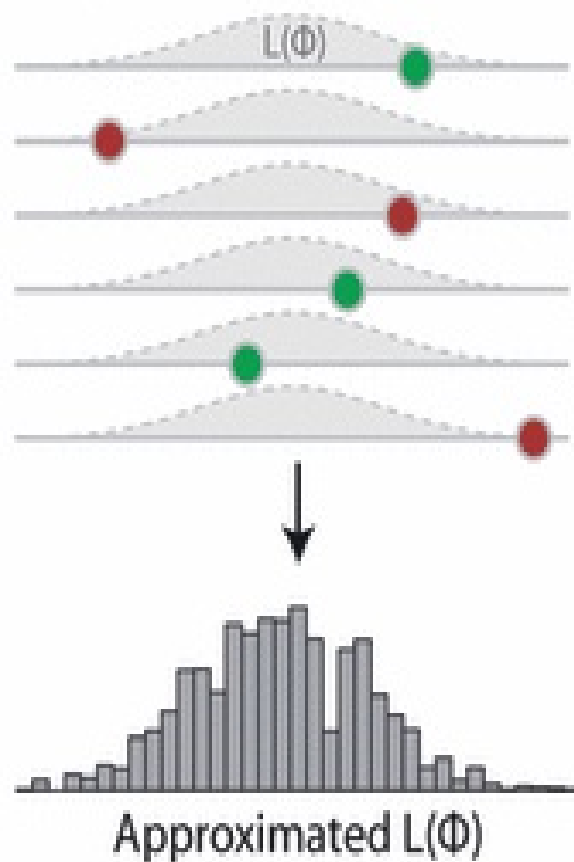
- Want to sample from some distribution  $g(x)$
- Requires that we can sample from a second distribution  $f(x)$  such that  $C \cdot f(x) > g(x)$  for all  $x$
- Algorithm
  - Draw a random value from  $f(x)$
  - Calculate the density  $g(x)$  and  $f(x)$  at that  $x$
  - Calculate  $a = g(x) / [C \cdot f(x)]$
  - Accept the proposed  $x$  with probability  $a$  based on a Bernoulli trial
  - If rejected, repeat by proposing a new  $x$ ...



# Sequential Monte Carlo (SMC)

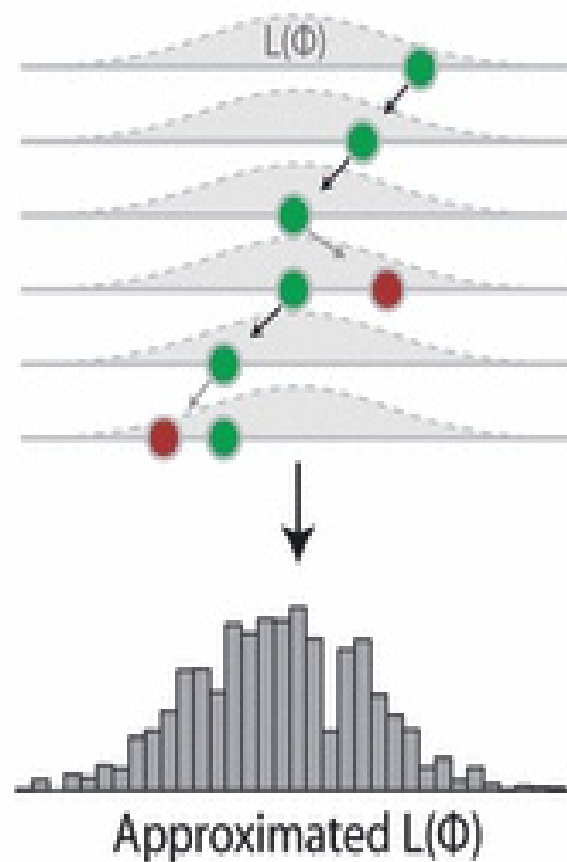
- Propose LARGE number of samples from prior
- Calculate Likelihood at each,  $L_i$
- Approximate normalizing constant  $P(Y) \propto \sum L_i$
- Calculate weights  $w = L_i/P(Y)$
- Resample proportional to weights (Inv CDF)
- Risks:
  - If  $n$  is small, weights concentrated
  - Harder in higher dimensions, broad priors
- Through time = Particle Filter

### Rejection Sampling (REJ)



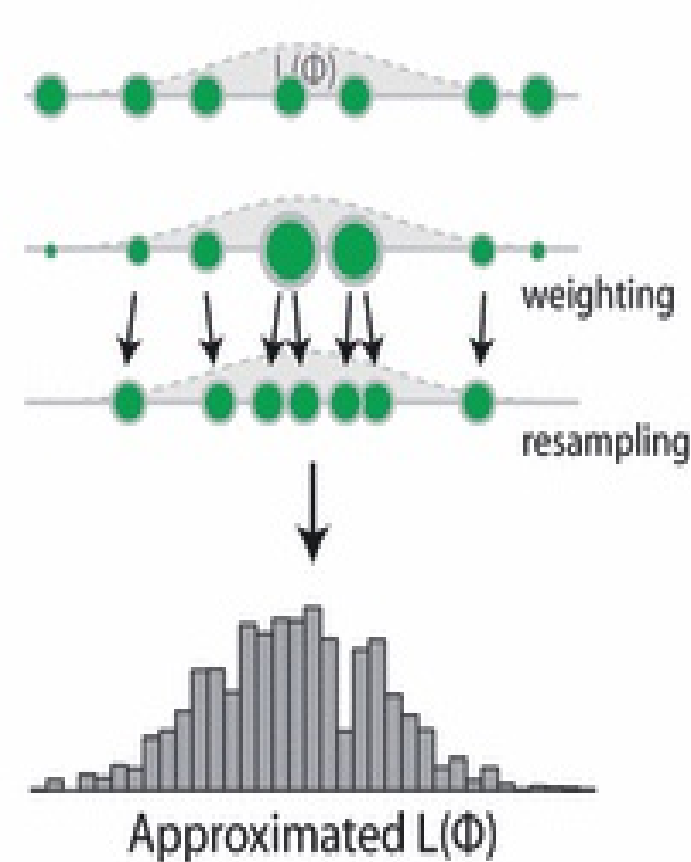
- 1) Draw a parameter  $\Phi$
- 2) Calculate  $L(\Phi)$
- 3) Accept proportional to  $L(\Phi)$

### MCMC Algorithm



- 1) Draw new parameter  $\Phi'$  close to the old  $\Phi$
- 2) Calculate  $L(\Phi')$
- 3) Jump proportional to  $L(\Phi')/L(\Phi)$

### SMC Algorithm



- 1) Last set of parameters  $\{\Phi_i\}$
- 2) Assign weight  $\omega_i$  proportional to  $L(\Phi_i)$
- 3) Draw new  $\{\Phi_i\}$  based on the  $\omega_i$